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Flight Mechanics Technical Memorandum 410

A FORTRAN PROGRAM FOR THE CALCULATION OF THE CALIBRATION COEFFICIENTS OF A SIX-COMPONENT STRAIN GAUGE BALANCE



by

S.S.W. Lam

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SUMMARY

A computer program written in FORTRAN has been developed to calculate the calibration coefficients of a six-component strain gauge balance. The method of Least Squares is used in the calculation procedures which do not require the components of the balance to be loaded independently. The procedures are, in fact, independent of the way the balance is loaded as long as the supplied load/output data provide sufficient information for the determination of the coefficients. The program incorporates several statistical estimations for the evaluation of the performance of the calibration process. It can therefore be used as an experimental tool for the investigation of different loading schemes which may be used to calibrate a strain gauge balance.



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Contents

Nomenclature

	Inti	roauct	ion	1		
2	Formulations of the Least Squares Method			3		
3	Statistical Estimations					
	3.1	Stand	ard Error of Estimates	5		
	3.2	Coeffi	cient of Multiple Correlation	5		
	3.3	Coeffi	cient of Partial Correlation	6		
	3.4	95% (Confidence Interval	7		
4	The Computer Program SGBCalib					
	4.1	4.1 Program Variables				
	4.2	Progr	am Structure	8		
		4.2.1	Program Inputs and Outputs	9		
		4.2.2	Evaluations of the Matrices	10		
		4.2.3	Statistical Estimations	10		
		4.2.4	Evaluation of Coefficient of Partial Correlations	11		
5	Pro	Program Listing •				
6	Example					
	Reference					

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Nomenclature

[A]	a matrix derived from the Least Squares criteria, defined by equation (7)
a_{ij}	coefficient of partial correlation
$C_{i.j}$	coefficient of calibration
[C], [C1], [C2]	calibration coefficient matrices
[E]	a matrix derived from the Least Squares criteria, defined by equation (5)
H_i	a component load of strain gauge balance
[H]	vector of component loads
$[H^*]$	vector of "non-linear" loading terms
N	total number of sets of data
N_f	degree of freedom in statistical estimates
p	index of summation
R_i	a component of strain gauge balance output
r_i	coefficient of multiple correlation
$S_{H,i}$	standard deviation of H_i
$S_{RH,i}$	standard error of estimates
X_i	a variable in regression equation

1 Introduction

In wind tunnel operations, calibrating a multi-component strain gauge balance is a process of determining the coefficients of an equation which describes the relationship between the applied loads and the strain gauge cutputs from each component. This involves choosing an appropriate calibration equation which best describes the behaviour of the particular balance. The balance is then loaded in a specific manner and the output from each strain gauge component is noted. A sufficient number of loadings has to be applied to the system, depending on the complexity of the chosen calibration equation, in order that the coefficients of the calibration equation may be totally determined.

Ideally, a perfect strain gauge balance would have each strain gauge output corresponding to one and only one load component. However, in a real situation, it is impossible to completely eliminate interactions from other components. Very often, significant non-linear interactions (functions of a combination of load components) are present as well as the usual linear ones (functions of a single load component). For a well designed 6-component balance, it is usual to express the strain gauge output of the *i*th balance component $(R_i, i = 1, ..., 6)$ as a second order polynomial function of all the components of the applied load $(H_j, j = 1, ..., 6)$. Thus the form of the calibration equation of a 6-component strain gauge balance is

$$R_{i} = C_{i,1}H_{1} + C_{i,2}H_{2} + \dots + C_{i,6}H_{6} + C_{i,11}H_{1}^{2} + C_{i,22}H_{2}^{2} + \dots + C_{i,66}H_{6}^{2} + C_{i,12}H_{1}H_{2} + C_{i,13}H_{1}H_{3} + \dots + C_{i,56}H_{5}H_{6}, \qquad i = 1, ..., 6$$
(1)

where the C's are the calibration coefficients determined during balance calibration and H_i are the component loads. The calibration coefficients of each component i may be classified as follows:

- 1. "linear", e.g. $C_{i,j}$ for j = 1, ..., 6;
- 2. "load squared", e.g. $C_{i,j}$ for j = 1, ..., 6;
- 3. "load cross product", e.g. $C_{i,jk}$ for $j=1,\ldots,5$, and $k=(j+1),\ldots,6$.

However, due to the difference in characteristics of a strain gauge in tension and in compression, an odd function representation is more appropriate to describe its behaviour on both sides of the origin. Hence the load/output relationship of a strain gauge balance may better be accounted for by the inclusion of the "load cubed" terms in the calibration equation. Thus the right hand side of equation (1) is modified with the addition of the terms

$$C_{i.111}H_1^3 + C_{i.222}H_2^3 + \cdots + C_{i.666}H_6^3$$
.

The set of six calibration equations for the six components can be represented in matrix form, with the calibration matrix partitioned into "linear" and "non-linear" sub-matrices:

$$|R| = |C1||H| + |C2||H^*| \tag{2}$$

where

[R] is the vector of strain gauge balance component outputs,

[C1] is the matrix of linear calibration coefficients, $C_{i,j}$,

[C2] is the matrix of non-linear calibration coefficients, $C_{i,jj}$, $C_{i,jk}$ and $C_{i,jjj}$,

[H] is the vector of component loads, and

 $[H^*]$ is the vector of squares, cross products and cubes of component loads.

The conventional method of calibrating a strain gauge balance of the type given by equation (2) is to apply pure component loads independently. When a single component is loaded, the outputs of all components of the strain gauge balance respond to one and only one load component. The "linear", "load squared" and the "load cubed" coefficients may hence be determined. The "load cross product" coefficients are found by applying two pure component loads simultaneously to the balance keeping one load constant at a finite value while varying the other. This method, although straightforward to use, relies on pure component loads which rarely represent the real, practical loading situation which the model encounters. It contravenes the basic principle that the loadings of the balance during the calibration should be representative of that experienced by the model during wind tunnel tests.

Theoretically, if the calibration equation describes exactly the relationship between the applied load and the strain gauge output, the way in which the balance is loaded during calibration should not affect the outcome of the results. Whether the balance components are loaded independently or simultaneously, the values of the calibration coefficients should be the same. However, in reality, the calibration equation given by equation (2) is only an approximation and does not describe exactly the load/output relationship of a practical strain gauge balance. The responses of the strain gauge balance components can be quite different under different loading distributions which causes different deflection characteristics in the balance. The calibration coefficients thus calculated are dependent of how the balance is loaded during the calibration.

Ramaswamy et al. [1] proposed a Least Squares (LS) method for the calibration of wind tunnel balances. This method does not require the balance components to be loaded independently. In fact, any combination of component loads may be applied simultaneously. The distributions of loads on the balance during the calibration can be chosen to match as closely as possible those likely to be experienced by the model during wind tunnel tests. In addition, there is no restriction of the sequence or pattern in which the loads have to be applied in order to obtain a particular calibration coefficient. Virtually any random set of component loads may be applied to the balance in any sequence provided that all the data so obtained generate enough independent information for the calculations of the coefficients. Another advantage of the LS method is that the "linear" and "non-linear" coefficients can be determined at the same time. This is because by loading all the components simultaneously, the "linear" and "non-linear" interactions among the components are all accounted for in the data. Less effort will hence be required to calibrate the balance if one can design an optimum loading scheme which would simulate the real loading situation as well as providing the most

information for the determination of the coefficients with the least number of sets of loads.

A major shortcoming of the LS method is that it assumes the independent variables, i.e. the H's in equation (1), are free of noises. In reality, due to certain factors such as friction or misalignment of the point of application of H_i , even though the same load is applied to the balance twice, the actual load experienced by the balance may not be exactly the same in both cases. Nevertheless, the errors arising from the difference between the load applied and the actual load experienced by the balance are usually insignificant when compared with the noises in the outputs R_i . Furthermore, for a well designed strain gauge balance, the noises in R_i are often satisfactorily small such that the LS method can normally be applied successfully.

This report describes a computer program, SGBCalib, which implements the LS method of determining the coefficients of the calibration equation given by (2). The only inputs to the program are the sets of applied component loads and the corresponding strain gauge outputs. If the given data contain sufficient information, the calibration coefficients will be evaluated. The program also generates useful statistical estimates of how closely the resulting calibration equation fits the given data. The reliability of the calibration equation is expressed in terms of the standard error of estimate of the strain gauge outputs for all the six components. The 95% confidence levels of the "linear" coefficients are also estimated as indications of the accuracy of the calibration.

Apart from the function of calculating the calibration coefficients from a given set of data, SGBCalib can be used to assist in the investigation of how the balance components should be loaded so as to yield the most information with the least number of sets of applied loads and at the same time best simulate the patterns of loading expected in the wind tunnel testings. It can also be used for studies of whether the errors of a particular strain gauge output could be reduced by further loadings designed to optimally minimize the uncertainties in the calibration coefficient.

2 Formulations of the Least Squares Method

Given that for a set of component loads, H_i where i = 1, ..., 6, the corresponding strain gauge balance outputs are R_i . If there are N sets of H_i with N sets of corresponding R_i , denoted by $H_{i,p}$ and $R_{i,p}$ where p = 1, ..., N, then the calibration coefficients may be found such that the sum of the squares of the difference between the measured strain gauge output and that obtained from the calibration equation is a minimum. That is

$$e_i = \sum_{p=1}^{N} \left[C_{i,1} H_{1,p} + C_{i,2} H_{2,p} + \dots + C_{i,666} H_{6,p}^3 - R_{i,p} \right]^2, \quad i = 1, \dots, 6,$$

should be a minimum. This involves partial differentiating each of the e_i with respect to each of the coefficients and equating the resulting expressions to zero. The result is a set of 33

non-linear simultaneous equations for each component i given as follows¹:

$$\sum \left[C_{i,1}H_1 + C_{i,2}H_2 + \dots + C_{i,666}H_6^3 - R_i \right] H_1 = 0$$

$$\sum \left[C_{i,1}H_1 + C_{i,2}H_2 + \dots + C_{i,666}H_6^3 - R_i \right] H_2 = 0$$

$$\vdots$$

$$\sum \left[C_{i,1}H_1 + C_{i,2}H_2 + \dots + C_{i,666}H_6^3 - R_i \right] H_6^3 = 0. \tag{3}$$

It is more convenient to put equations (3) in matrix form:

$$[E][C_i] = [A_i], i = 1, ..., 6,$$
 (4)

where

$$[E] = \begin{bmatrix} \sum H_1 H_1 & \sum H_1 H_2 & \cdots & \sum H_1 H_6^3 \\ \sum H_2 H_1 & \sum H_2 H_2 & \cdots & \sum H_2 H_6^3 \\ \vdots & \vdots & & \vdots \\ \sum H_6^3 H_1 & \sum H_6^3 H_2 & \cdots & \sum H_6^3 H_6^3 \end{bmatrix}$$
 (5)

is a 33 \times 33 square matrix and is common for all the six components. The two column matrices, $[C_i]$ and $[A_i]$, for the *i*th component are given by

$$[C_i] = \begin{bmatrix} C_{i,1} \\ C_{i,2} \\ \vdots \\ C_{i,666} \end{bmatrix} \qquad [A_i] = \begin{bmatrix} \sum_{i=1}^{n} H_1 R_i \\ \sum_{i=1}^{n} H_2 R_i \\ \vdots \\ \sum_{i=1}^{n} H_6^3 R_i \end{bmatrix}.$$

The $[C_i]$'s and $[A_i]$'s for all the six components may be combined to form two 33×6 matrices, [C] and [A], such that

$$[C] = \begin{bmatrix} C_{1.1} & C_{2.1} & \cdots & C_{6.1} \\ C_{1.2} & C_{2.2} & \cdots & C_{6.2} \\ \vdots & \vdots & & \vdots \\ C_{1.666} & C_{2.666} & \cdots & C_{6.666} \end{bmatrix},$$

$$(6)$$

$$[A] = \begin{bmatrix} \sum H_1 R_1 & \sum H_1 R_2 & \cdots & \sum H_1 R_6 \\ \sum H_2 R_1 & \sum H_2 R_2 & \cdots & \sum H_2 R_6 \\ \vdots & \vdots & & \vdots \\ \sum H_6^3 R_1 & \sum H_6^3 R_2 & \cdots & \sum H_6^3 R_6 \end{bmatrix}$$
(7)

⁴ For the sake of clarity, the subscript p for R_t and H_t is omitted in the remainder of the report. Whenever the summation symbol is used, it is understood that the sum is to be taken over p = 1, ..., N.

The Least Squares criteria can therefore be represented by the matrix equation

$$[E][C] = \{A\}. \tag{8}$$

Provided that [E] is non-singular, the coefficient matrix [C] may be found as

$$[C] = [E]^{-1}[A]. (9)$$

3 Statistical Estimations

In this section we adopt a few statistics to measure the accuracy and reliability of the calibration of the balance using the LS method. We seek to determine how well the resulting calibrating equation describes the relationship between the strain gauge outputs and the applied component loads. The statistics described below provide us with information on and measurements of the goodness of fit of the calibration equation, and the uncertainties of the coefficients determined by the LS method.

3.1 Standard Error of Estimates

The standard error of estimates of the ith component is given by

$$S_{RH,i} = \sqrt{\frac{\sum (R_i - \hat{R}_i)^2}{N_f}} \tag{10}$$

where R_i is the measured ith component of the strain gauge balance output.

 R_i is the corresponding estimated value using the calibration equation.

 N_f is the degree of freedom, which is (N-33), and

N is the total number of sets of loading cases.

It is a measure of the scatter of the balance outputs about the estimated values. It has properties analogous to those of the standard deviation, i.e., if N is large enough, there would be approximately 68%, 95% and 99.7% of the measured R_i lies within the bounds of $S_{RH,i}$, $2S_{RH,i}$ and $3S_{RH,i}$ respectively from the estimated \hat{R}_i value.

3.2 Coefficient of Multiple Correlation

The coefficient of multiple correlation of the *i*th component, r_i , is defined as

$$r_i = \sqrt{\frac{\sum (\hat{R}_i - R_i)^2}{\sum (\hat{R}_i - R_i)^2 + \sum (R_i - \hat{R}_i)^2}}$$
(11)

The term $\sum (\hat{R}_i - \hat{R}_i)^2$ is known as the explained variation, so called because it accounts for the variations (of R_i from the mean) which are explained by the regression. $\sum (R_i - \hat{R}_i)^2$ is the unexplained variation. It represents the variations which remain unexplained after the regression. The coefficient of multiple correlation shows what fraction of the overall variations is accounted for by applying the calibration equation. It is therefore a measure of the correlation between a balance component output and all of the component loads. It shows how well the calibration equation describes the relationship between the outputs of the strain gauge balance and the component loads. It is also a measure of the closeness of fit of the calibration equation to the actual measured strain gauge balance outputs. The value of r_i lies between 0 and 1. The closer it is to 1 the better is the linear relationship. When $r_i = 1$ the correlation is called perfect.

3.3 Coefficient of Partial Correlation

The coefficient of partial correlation measures the linear relationship between a component of the strain gauge balance output and a single component load while the other component loads are kept constant.

The linear correlation coefficients a_{12} between two variables X_1 and X_2 is given by

$$a_{12} = \frac{N \sum X_1 X_2 - (\sum X_1)(\sum X_2)}{\sqrt{[N \sum X_1^2 - (\sum X_1)^2][N \sum X_2^2 - (\sum X_2)^2]}}$$
(12)

where N is the number of data points. a_{12} is called the zero order correlation coefficient. We can denote the coefficient of partial correlation between X_1 and X_2 keeping X_3 constant by $a_{12,3}$ (the 1st order correlation coefficient) such that

$$a_{12.3} = -\frac{a_{12} - a_{13}a_{23}}{\sqrt{(1 - a_{13}^2)(1 - a_{23}^2)}}.$$
 (13)

Similarly the coefficient of partial correlation between X_1 and X_2 keeping X_3 and X_4 constant may be given by (see for e.g. Spiegel [2])

$$\frac{a_{12,34}}{\sqrt{(1-a_{13,4}^2)(1-a_{23,4}^2)}} = \frac{a_{12,3}-a_{14,3}a_{24,3}}{\sqrt{(1-a_{13,4}^2)(1-a_{23,4}^2)}} = \sqrt{(1-a_{14,3}^2)(1-a_{24,3}^2)}.$$
(14)

The coefficient of partial correlation between the strain gauge balance output R_i and the component load H_1 , say, keeping the other component loads constant may therefore be denoted by $a_{B_i1/23456}$, where the numerical subscripts denote the component loads, and can be deduced from equations (12), (13) and (14).

An ideal strain gauge balance would have perfect linear correlation (a-1) between an output and its corresponding component load and zero correlation with the rest of the components. These coefficients therefore indicate how far the strain gauge balance has fallen short of this ideal behaviour

3.4 95% Confidence Interval

To establish a 95% confidence interval for a calibration coefficient $C_{i,j}$ we can construct the t-distribution statistics:

$$t = \frac{C_{i,j} - \bar{C}_{i,j}}{S_{RH,i}/S_{H,i}} \sqrt{N_f}$$
 $i = 1, ..., 6, j = 1, ..., 6$

where $S_{RH,i}$ is the standard error of estimate of the *i*th component given in section 3.1, $S_{H,j}$ is the standard deviation of H_j , $\ddot{C}_{i,j}$ is the expected mean of $C_{i,j}$.

The 95% confidence interval for the coefficient $C_{i,j}$ is then given by the expression

$$\frac{t_{0.975}}{N_f} \left(\frac{S_{RH,i}}{S_{H,j}} \right).$$

The percentile values $t_{0.975}$ for various degree of freedom (N_f) may be found in most reference books in statistics (see e.g. Yamane [3]).

4 The Computer Program SGBCalib

The program SGBCalib was written in FORTRAN and was implemented in DEC¹ MicroVAX¹ II operating under the VMS¹ (Version 4.5) system. The program, however, was not intended to be VMS-specific, and the language used conforms to the American National Standard FORTRAN-77 (ANSI X3.9-1978).

4.1 Program Variables

Drag measurements using wind tunnel force balance usually require high order of accuracy. In order to maintain data integrity, all the program variables concerned are assigned as "double precision". Most of the variables were passed between the various subroutines via common blocks:

/matrixA/, /matrixC/, /matrixE/, /inv/ These common blocks contain the elements of the matrices [A], [C], [E] and $[E]^{-1}$ respectively.

/loads/,/readings/

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They store all the given H_i and R_i data in the 2-D arrays HH and RR respectively. The program at present allows a maximum of 200 sets of H_i and R_i values to be processed. This may be changed if necessary by altering the dimension of the second subscripts of the statement

double precision HH(6,200), RR(6,200).

The dimension of the second subscript of the variable res(6,200) in the subroutines Statistics, SumStats and Chart should also be changed to be the same as those of HH and RR.

/Hvector/ This is an array of the 33 H terms in the calibration equation. That is, it contains the values of

$$H_1, H_2, \ldots, H_6, H_1^2, H_2^2, \ldots, H_6^2, H_1H_2, H_1H_3, \ldots, H_5H_6, H_1^3, H_2^3, \ldots, H_6^3.$$

/residual/ This common block appears in the statistical routines only. It consists of the residual values of R_i , i.e. $(R_i - \hat{R}_i)$. The block also contains the maximum value of the residual, resmax, which is used for scaling of the chart of the residuals.

/misc/ This is a miscellaneous block containing the pth set of values of R_i , \hat{R}_i and H_i . The mean values of R_i are also included in this block.

/statsvar/ This block contains the variables used for the calculations of the various statistical estimates for the calibration. The major variables are:

$$\begin{array}{lll} \text{sRH(i,j)} & = & \sum \hat{R}_i H_j, & i,j=1,...,6 \\ \text{sR(i)} & = & \sum \hat{R}_i \\ \text{sR2(i)} & = & \sum \hat{R}_i^2 \\ \text{sH(i)} & = & \sum H_i \\ \text{sH2(i)} & = & \sum H_i^2 \\ \text{sHH(i,j)} & = & \sum H_i H_j, & i,j=1,...,6 \\ \text{dRRe2(i)} & = & \sum (R_i - \hat{R}_i)^2 \\ \text{dRRm2(i)} & = & \sum (\hat{R}_i - \bar{R}_i)^2 \\ \text{dReRm2(i)} & = & \sum (\hat{R}_i - \bar{R}_i)^2 \end{array}$$

4.2 Program Structure

The program is made up of different subroutines designated for various specific tasks. The listing of these subroutines is given in Section 5. The tasks undertaken by these subroutines may be divided broadly into three areas:

1. Program inputs and outputs.

- 2. Calculations of the matrices [A], [E], $[E]^{-1}$ and [C].
- 3. Estimation of the statistics of the calibration.

Apart from the various subroutines written specifically for the LS method of calculating the calibration coefficients, there are two main subroutines which are used in matrix operations. They are the matrix inversion (MatInv) and the matrix multiplication (MatMult) routines. For the sake of completion, these subroutines are also listed in Section 5.

4.2.1 Program Inputs and Outputs

The only program inputs are the sets of H_i and R_i values which are read from the data file SGBIN.DAT. The format of the file is as follows:

- 1. The first line is empty and is skipped by the program. It may be used for data identification purpose.
- 2. The second line contains a row of 6 components of loads, H_i , each of which is separated from the next value by one or more spaces.
- 3. The third line is a row of 6 components of strain gauge outputs, R_i , corresponding to each of the H_i value in the above row.
- 4. This pattern: an empty line, a row of 6 H_i values followed by a row of 6 corresponding R_i values, repeats throughout the data file for as many sets of loading as are available.

When the program is executed, all the relevant information and results are written of an output file SGBOUT. DAT which may subsequently be printed on a line printer. The output of the program consist of the following:

- 1. All the values of H_i and R_i provided in the input file.
- 2. The evaluated calibration coefficient matrix [C] in the form:

- 3. Statistical estimations which include:
 - the standard errors of estimates for each of the component,
 - the coefficients of multiple correlation,
 - the coefficients of partial correlation between R_i and each of the component loads,
 - the 95% confidence level of the linear coefficients, i.e.

- 4. A table of comparisons of the measured and estimated values of R_i and the residuals.
- 5. A line printer plot of the residuals of R_i over \hat{R}_i for each of the 6 components. The maximum deviation is ± 10 characters from the zero marker (the character "!") and is scaled to be equivalent to the maximum residual value of all 6 components.

4.2.2 Evaluations of the Matrices

As each set of the H_i and R_i values are read from the input file, the "load squares" (H_i^2) , "load cubes" (H_i^3) and "load cross products" $(H_iH_j, i = 1, ..., 5 \text{ and } j = i + 1, ..., 6)$ are evaluated by means of the subroutine Hfactors. The elements forming the matrices [E] and [A] are then calculated and summed in the subroutine SumElements.

After all the H_i and R_i values have been read, $[E]^{-1}$ is found using the matrix inversion routine MatInv. An error flag (designated by the variable error) is used to indicate if an inverse has successfully been calculated. When error = 0, the inverse is found and no error occurs. When error = 2, [E] does not have an inverse and the program terminates with the message

%%% INVERSE OF [E] DOES NOT EXIST %%%

appearing on the terminal screen. When $[E]^{-1}$ is found the coefficient matrix [C] is evaluated according to equation (9) by the matrix multiplication routine MatMult and the elements of [C] are written to the output file SGBOUT.DAT.

4.2.3 Statistical Estimations

The evaluations of the various statistics mentioned in Section 3 are initiated by the subroutine Statistics. The estimated value of R_i , denoted as Re(i) in the program, is calculated

according to equation (2). Before equation (2) can be applied, the coefficient matrix [C] is transposed and partitioned into a linear matrix [C1] and a non-linear matrix [C2]. This is accomplished by the statements

```
do i = 1, 6
  do j = 1, 6
    c1(i,j) = c(j,i)
  end do
  do j = 1, 27
    c2(i,j) = c(j+6,i)
  end do
end do
```

in the Statistics routine. The subroutine SGBReading then evaluates Re(i) from the data of H_i according to equation (2) and the various summations required by the statistics are calculated in the subroutine SumStats.

The various statistics described in Section 3 are designated as

```
S(i) for the standard error of estimates S_{RH,i}, cmc(i) for the coefficients of multiple correlation r_i, and cl(i,j) for the 95% confidence levels of the linear coefficients C_{i,j}.
```

Because of the complexity of the calculations of the coefficients of partial correlation, they are described in more detail in Section 4.2.4. The calculation of the confidence levels requires the 0.975 percentile values at various degree of freedom for the t-statistics. These values are provided as a look-up table in the subroutine Tstats.

A graphical representation of the residuals of R_i over \hat{R}_i is constructed by the subroutine Chart. The results for all six components are displayed together across the width of the page. The zero axes are represented by the character "!". The residuals for components 1 and 4 are represented by the character "*", those for components 2 and 5 by the character "o" and those for components 3 and 6 by the character "#". The scale of the deviation is based on the maximum value of the residual (designated by the variable resmax in the program) of all six components. The value of resmax occupies a space of 10 character on the chart and the other values of the residual are scaled accordingly.

4.2.4 Evaluation of Coefficient of Partial Correlations

For the sake of clarity, let us consider a single component of the strain gauge balance output R_3 , say, and seek to determine the correlation coefficients between R_3 and each of the component loads H_i keeping the other component loads constant. It is convenient, then, to consider

an array X_i , $i=1,\ldots,7$ such that $X_1=R_3$, $X_2=H_1$, $X_3=H_2$, ..., and $X_7=H_6$. The coefficients of partial correlation between R_3 and each of the H's may therefore be denoted as

 $a_{12.34567}$, $a_{13.24567}$, $a_{14.23567}$, $a_{15.23467}$, $a_{16.23457}$, and $a_{17.23456}$.

To determine, say, $a_{12.34567}$, we may use the relationship

$$\frac{\vec{a}_{12.34567}}{\vec{a}_{12.34567}} = \frac{a_{12.4567} - a_{13.4567} a_{23.4567}}{\sqrt{(1 - a_{13.4567}^2)(1 - a_{23.4567}^2)}} \tag{15}$$

in which the variable X_3 is used as the "mediator". Theoretically, we can use any of the "kept-constant" variables X_3 , X_4 , X_5 , X_6 and X_7 to be the mediator to calculate $a_{12.34567}$. For convenience in programming, we opt for the lowest subscript of the "kept-constant" variables.

To determine the next lower order correlation coefficient, say, $a_{13.4567}$ in the right hand side of equation (15), we, therefore use the relationship

$$a_{13.4567} = \frac{a_{13.567} - a_{14.567}a_{34.567}}{\sqrt{(1 - a_{14.567}^2)(1 - a_{34.567}^2)}} \tag{16}$$

in which X_4 is now the mediator. Ultimately, these coefficients are related to the zero order correlation coefficients a_{12} , a_{13} , ..., a_{23} , a_{24} , ..., a_{67} where

$$a_{ij} = \frac{N \sum X_i X_j - (\sum X_i)(\sum X_j)}{\sqrt{[N \sum X_i^2 - (\sum X_i)^2][N \sum X_j^2 - (\sum X_j)^2]}}$$
(17)

In the program, the determination of the coefficients of partial correlation is carried out by the subroutine PartCorr. The sums $\sum X_i$, $\sum X_i^2$, and $\sum X_i X_j$ are mapped into the arrays sumX(i), sumX2(i), and sumXX(i) respectively. The correlation coefficients are denoted by the 3-D array corr(m,i,j) in which (m-1) is the order of the correlation, i is the subscript of the dependent variable X_i and j is that of the independent variable X_j . The values of the correlation that we are looking for are therefore stored in corr(6,1,k), k = 2,...,7.

The zero order correlation coefficients are evaluated first and are determined by the statements

The mediator for the determination of each order of the correlation coefficients is preselected in the subroutine Rank. The "ranking" of the mediator subscript is stored in the array mediator(m) so that the highest order (i.e. m = 6) would have the lowest subscript in the range of 2 to 7 excluding the value of k when determining corr(6,1,k). The next highest order would then have the next lowest subscript and so on. For example, to determine the coefficient $a_{15.23467}$, the "ranking" of the mediator subscripts is

order : 2 3 4 5 6 mediator : 7 6 4 3 2

To calculate the 1st order correlation coefficients, the variable with the subscript 7 is to be the mediator. Hence the coefficients $a_{12.7}$, $a_{13.7}$, $a_{14.7}$, ..., $a_{56.7}$ are evaluated.

Because of the symmetrical relationship that corr(m,i,j) = corr(m,j,i), only the values of corr(m,i,j) in which i < j are evaluated and used in the calculations.

5 Program Listing

The following pages are a full listing of the program SGBCalib. The program is intended to be executed interactively on a computer terminal. It may also be submitted as batch job in most system without any modification.

The program listing is roughly divided into 3 sections:

- Main section This section consists of the main routine SGBCalib and the subroutines SumElements and Hfactors. They are the routines that implement the LS method.
- Statistical section This includes the subroutines Statistics, StatsInitialize, SumStats, SGBReading, PartCorr, Rank, Chart and Resetline and are responsible for the evaluations of the statistical variables.
- Matrix Operations The matrix inversion MatInv and matrix multiplication MatMult routines are included for the sake of completion of the program listing. There are other minor subroutines following MatInv and are called by MatInv only.

```
Program SGBCalib
   Strain Gauge Balance Calibration using Least Squares Method
    Given a set of S.G.B. readings, R(i), from a set of
    known loadings, H(i), the matrices [E] and [A] are formed.
    The coefficient matrix, [C], is obtained according to
     the equation
            [C] = [Einv][A]
   Written by S. Lam - 31 March 1988
   Associated routines: - Hfactors(H)
                        SumElements(R)
                        Statistics(N)
                        Hfactors(H)
                        MatInv(dim, E, Einv, error)
                        MatMult(C1,m,n,C2,n,k,C3,error)
     double precision E(33,33), Einv(33,33), A(33,6), C(33,6)
     double precision H(6),R(6),Hstar(27),Rm(6)
     double precision sumR(6), HH(6,200), RR(6,200)
     integer error, Hsubscript(33)
     common /inv/Einv /matrixE/E /matrixA/A /Hvector/H, Hstar
     common /matrixC/C /loads/HH /readings/RR /misc/Rm
     data Hsubscript/1,2,3,4,5,6,11,22,33,44,55,66,12,13,14,15,16,
          23,24,25,26,34,35,36,45,46,56,111,222,333,444,555,666/
   Files for I/0:-
С
       SGBIN.DAT = file containing sets of H(i) and R(i) for
С
                    data input
       SGBOUT.DAT = file for output of results
     open(unit=3, file='SGBIN.DAT', status='old')
     open(unit=4, file='SGBOUT.DAT', status='new')
C -----
C Print Headings of the program outputs
     write(4,10)
  10 format(/,' STRAIN GAUGE BALANCE CALIBRATION USING LEAST'
        ' SQUARES METHOD'/' -----'
        '----'//' The Supplied Component Loads'
```

' and Strain Gauge Bridge Readings:-'/)

```
write(4,15) (i, i=1,6), (j, j=1,6)
 15 format(6(5x, 'H', i1, 3x), 6(5x, 'R', i1, 3x))
  Initialize the [E] and [A] matrices
C -----
     N = 0
     do n_row = 1,33
       do n_col = 1,33
        E(n_row, n_col) = 0.0
       end do
     end do
     do n_col = 1,6
       do n_row = 1.33
        A(n_row, n_col) = 0.0
       end do
       sumR(i) = 0.0
     end do
C -----
C Input data :- H(i) = SGB readings vector
              R(i) = Load vector
 20 read(3,*,end=50)
     read(3,*,end=50) (H(i),i=1,6)
     read(3,*,end=50) (R(i), i=1,6)
     write(4,30) (H(i), i=1,6), (R(i), i=1,6)
 30 format(12(x,f9.5))
     N = N + 1
     do i = 1,6
       sumR(i) = sumR(i) + R(i)
C -----
C Store the original H(i) and R(i) values for later computations
       HH(i,N) = H(i)
       RR(i,N) = R(i)
C Evaluate the Least Squares elements for the [E] and [A] matrices
     call Hfactors(H)
     call SumElements(R)
     go to 20
 50 continue
     write(4,60) N
  60 format('
                Number of sets of loading = ', I3,/)
```

```
C ------
C Evaluate the inverse of [E]
C ------
    call MatInv(33, E, Einv, error)
    if (erro .gt. 0) go to 80 !{ No inverse }
C -----
C Evaluate the coefficient matrix [C] = [Einv][A]
C -----
    call MatMult(Einv,33,33,A,33,6,C,error)
    write(4,70) (i, i=1,6)
 70 format(1H1,/,' The calibration coefficient matrix [C]:-'//,
    *4x,' H''s',6(7x,'C',i1,'-',4x))
    do i = 1,33
      write(4,100) Hsubscript(i), (C(i,j), j=1,6)
    end do
    do i = 1,6
      Rm(i) = sumR(i)/float(N) !{ mean values of R(i) }
 call Statistics(N)
    go to 110
 80 write(*,90)
 90 format(/' %%% INVERSE OF [E] DOES NOT EXIST %%%')
100 format(4x,i3,6(2x,f12.7))
110 close(4)
    end
```

```
Subroutine SumElements(R)
C -----
   Evaluate the sums which make up the elements
  in the [E] and [A] matrices
     double precision A(33,6), E(33,33), Hfactor(33), R(6)
     common /matrixE/E /matrixA/A /inv/Einv /Hvector/Hfactor
     integer row, col
     do row = 1,33
       do col = 1,33
         E(row,col) = E(row,col) + Hfactor(row)*Hfactor(col)
       end do
     end do
     do col = 1,6
       do \ low = 1,33
         A(row,col) = A(row,col) + R(col)*Hfactor(row)
     end do
     return
     end
```

```
Subroutine Hfactors(H)
  Calculate the non-linear load factors
c -----
     double precision H(6), Hfactor(33)
     common /Hvector/Hfactor
     do i \approx 1,6
       Hfactor(i) = H(i)
                                    !{ linear }
       Hfactor(i+6) = H(i)*H(i)
                                   !{ squared }
       Hfactor(i+27) = H(i)*H(i)*H(i) !{ cubed }
     end do
     i = 13
     do j = 1,5
       do k = (j+1),6
         Hfactor(i) = H(j)*H(k)
                                !{ cross product }
         i = i+1
       end do
     end do
     return
     end
```

```
Subroutine Statistics(N)
C This subroutine evaluates the various statistics relating to
{\tt C} the {\tt R(i)} values as estimated by the calibration equation
С
C
     Input: N = number of sets of loading
C
C
     Associate routines :- SumStats(j)
С
                         StatsInitialize
C
                         SGBReading(C1,C2,Re,H,error)
С
                         Tstats(ndf,t)
C
                         PartCorr(N,i,cpc)
C
                         Chart(scale,N)
     double precision C(33,6),C1(6,6),C2(6,27),R(6),H(6),Re(6),Rm(6)
     double precision HH(6,200), RR(6,200)
     double precision sumRH(6,6),sumR(6),sumH(6),sumR2(6),sumH2(6)
     double precision sumHH(7,7), Se(6), SdR(6), SdH(6), cmc(6), cl(6,6)
     double precision cpc(6,6),dRRm2(6),dReRm2(6),dRRe2(6),res(6,200)
      common /Sums/sumRH, sumR, sumH, sumR2, sumHH,
                   dRRe2,dReRm2,dRRm2
      common /loads/HH /readings/RR /misc/Rm,R,Re,H
      common /residual/res,resmax /matrixC/C
     parameter (nv=33)
                         !{ no. of variables }
     resmax = 1.0e-20
                             !{ maximum res (residual) for chart scaling }
     call StatsInitialize !{ Initialize the variables to zero }
  Break up the coefficient matrix into
  linear (C1) and non-linear (C2) matrices
C -----
     do i = 1, 6
       do j = 1,6
         C1(i,j) = C(j,i)
       end do
       do j = 1,27
         C2(i,j) = C(j+6,i)
       end do
       C Restore the original H(i) and R(i) values for computations
     do j = 1, N
       do i = 1, 6
         H(i) = HH(i,j)
         R(i) = RR(i,j)
       end do
```

```
C Estimate the values of R, Re(i), from the coefficients matrices
C C1 and C2, and evaluate the various sums for statistical calculations
C -----
      call SGBReading(C1,C2,Re,H,error)
      call SumStats(j)
C Use HH(i,j) as storage for Re(i) for later print-outs
C -----
      do i = 1, 6
        HH(i,j) = Re(i)
      end do
    end do
 20 ndf = float(N-nv)
                     !{ degree of freedom }
    if (N .le. nv) ndf = 1
    fN = float(N)
    call Tstats(ndf,t) !{ find the t-statistics }
C ------
C Evaluate the statistics
   Se(i) = standard error of estimates
   SdR(i) = standard deviation of R(i)
   SdH(i) = standard deviation of H(i)
    cmc(i) = coefficient of multiple correlation
   cpc(ij) = coefficient of partial correlation
   cl(ij) = 95% confidence level of the linear coefficients
    do i = 1.6
      Se(i) = sqrt(dRRe2(i)/ndf)
      SdR(i) = sqrt(dRRm2(i)/ndf)
      SdH(i) = sqrt((sumH2(i)-sumH(i)*sumH(i)/fN)/ndf)
      cmc(i) = sqrt(dReRm2(i)/(dRRe2(i)+dReRm2(i)))
      call PartCorr(N,i,cpc)
    end do
    do j = 1,6
      do i = 1,6
        cl(i,j) = t/sqrt(float(ndf))*Se(i)/SdH(j)
      end do
    end do
C Write the results to the output file
write(4,30)
    write(4,110) (i, i=1,6)
    write(4,60) (Se(i), i=1,6)
 30 format(1H1,///,' Standard errors of estimates: ',/)
    write(4,40)
    write(4,110) (i, i=1,6)
```

```
write(4,60) (cmc(i), i=1,6)
 40 format(///, 'Coefficients of multiple correlation:-'/)
     write(4,50)
     write(4,120) (i, i=1,6)
 50 format(///, 'Coefficients of partial correlation:-'/)
     do j = 1, 6
      write(4,65) j, (cpc(i,j), i = 1,6)
     end do
 60 format(6(2x,f11.5))
 65 format(3x, 'H', i1,6(2x,f11.5))
     write(4,70) (i, i=1,6)
 70 format(/// )5% confidence level of the linear coeff.:-'//
            2x,6 5x,'C',i1,'-',4x))
     do j = 1, 6
      write(4,80) (cl(i,j), i = 1,6)
     end do
 80 format(6(2x,f10.7))
C -----
C Print out the measured , estimated and the residuals of R(i)
C ------
     write(4,90) (i, i=1,6)
 90 format(1H1,' Table of residuals:-',//,
    * 6(8x,'R',i1,11x)/6(' meas. est. res. '))
     do j = 1, N
      write(4,100) (RR(i,j),HH(i,j),res(i,j), i=1,6)
 100
      format(5(2(x,f5.3),x,f6.4,x,'|'),2(x,f5.3),x,f6.4)
     end do
 110 format(3x,6(6x,'R',i1,5x))
120 format(8x,6(6x,'R',i1,5x))
C -----
{\tt C}   
Construct the chart of residuals
C -----
     scale = 10.0/resmax
     call Chart(scale,N)
     return
     end
```

```
Subroutine StatsInitialize
```

```
c -----
C Initializes the variables of the Statistics routine to zero
     double precision sumRH(6,6),sumR(6),sumH(6),sumR2(6),sumH2(6)
     double precision sumHH(7,7),dRRe2(6),dRRm2(6),dReRm2(6)
     common /Sums/sumRH,sumR,sumH,sumR2,sumHH,sumR1
                  dRRe2,dReRm2,dRRm2
     do i = 1, 6
       do j = 1,6
         sumRH(j,i) = 0.0
         sumHH(j+1,i+1) = 0.0
       end do
       dRRe2(i) = 0.0
       dReRm2(i) = 0.0
       dRRm2(i) = 0.0
       sumR(i) = 0.0
       sumH(i) = 0.0
       sumR2(i) = 0.0
       sumH2(i) = 0.0
     end do
     return
     end
```

```
Subroutine SGBReading(C1,C2,R,H,error)
  This subroutine evaluates the six components of Strain Gauge
C Balance Readings (R) from the six components of Loads (H).
C Input: C1,C2 - linear and non-linear coefficient matrices
C Output: R (the SGB Reaings vector)
C Error: if error > 0, an error has occurred
C Associate routines:- Hfactors(H)
                     MatMult(C1,m,n,C2,n,k,C3,error)
     double precision C1(6,6),C2(6,27),X(6),R(6),H(6),Hstar(27)
     integer error
     common /Hvector/X, Hstar
     call Hfactors(H)
C -----
C Perform the matrix multipications:
     [C1][H] and [C2][H*]
     Call MatMult(C1,6,6,H,6,1,R,error)
     if (error .gt. 0) return
     Call MatMult(C2,6,27,Hstar,27,1,X,error)
     if (error .gt. 0) return
     do i = 1,6
       R(i) = R(i) + X(i)
      end do
      return
      end
```

```
Subroutine SumStats(j)
C Evaluate the sums for the statistical estimations
      double precision sumRH(6,6),sumR(6),sumH(6),sumR2(6),sumH2(6)
      double precision sumHH(7,7),res(6,200),dRRe2(6),dRRm2(6),dReRm2(6)
      double precision H(6),R(6),Re(6),Rm(6)
      common /Sums/sumRH,sumR,sumH,sumR2,sumH2,sumHH,
                    dRRe2,dReRm2,dRRm2
      common /misc/Rm,R,Re,H /residual/res,resmax
     do i = 1,6
        res(i,j) = R(i)-Re(i)
        if (abs(res(i,j)) .gt. resmax) resmax = abs(res(i,j))
        dRRe2(i) = dRRe2(i) + res(i,j)**2
        dRRm2(i) = dRRm2(i) + (R(i)-Rm(i))**2
        dReRm2(i) = dReRm2(i) + (Re(i)-Rm(i))**2
        sumR(i) = sumR(i) + R(i)
        sumH(i) = sumH(i) + H(i)
        sumR2(i) = sumR2(i) + R(i)*R(i)
        sumH2(i) = sumH2(i) + H(i)*H(i)
        do m = 1,6
          sumRH(i,m) = sumRH(i,m) + R(i)*H(m)
          sumHH(i+1,m+1) = sumHH(i+1,m+1) + H(i)*H(m)
        end do
      end do
      return
      end
```

```
Subroutine Tstats(ndf,t)
C -----
C Estimate the t-statistics from the degree of freedom
     dimension tdf(30)
     data tdf/12.706,4.303,3.182,2.776,2.571,2.447,2.365,2.306,
               2.262,2.228,2.201,2.179,2.160,2.145,2.131,2.120,
               2.110,2.101,2.093,2.086,2.080,2.074,2.069,2.064,
               2.060,2.056,2.052,2.048,2.045,2.042/
     if (ndf .le. 30) then
       t = tdf(ndf)
     else if (ndf .le. 40) then
       t = tdf(30)-0.021*float((ndf-30)/10)
     else if (ndf .le. 60) then
       t = 2.021-0.021*float((ndf-40)/10)
     else if (ndf .le. 120) then
       t = 2.000-0.020*float((ndf-60)/10)
     else
       t = 1.96
     end if
     return
     end
```

```
Subroutine PartCorr(N,iR,cpc)
C ------
C Subroutine to evaluate the coefficients of partial correlation
С
     Inputs: - N = total number of sets of data
С
             iR = the subscript of R whose coeff. of part. corr.
C
                  are to be evaluated
С
    Outputs: - cpc = arrays containing the coeff. of part. corr.
С
С
    Cautions: - Errors may occur if, during the evaluation process,
С
               the values of certain correlation coefficients are
С
               close to 1.
C
    Associated routine :- Rank(m,n)
     double precision sumRH(6,6),sumH(6),sumH(6),sumH2(6)
     double precision sumXX(7,7),sumX(7),sumX2(7)
     double precision corr(6,6,7),cpc(6,6),a,b,c
      integer mediator(6), order
      common /Sums/sumRH, sumH, sumH2, sumH2, sumXX
C Map the sums of R(iR) and H(i) into the array sumX(j)
      the sums of squares into the array sumX2(j), and
      the sums of cross products into sumXX(j,k)
     sumX(1) = sumR(iR)
     sumX2(1) = sumR2(iR)
     do j = 2,7
       sumX(j) = sumH(j-1)
       sumX2(j) = sumH2(j-1)
       sumXX(1,j) = sumRH(iR,j-1)
      end do
C Evaluate the zero order (order = 1) correlation coefficients first
     do m1 = 1.6
       do m2 = (m1+1),7
           corr(1,m1,m2) = (N*sumXX(m1,m2)-sumX(m1)*sumX(m2))
               /sqrt((N*sumX2(m1)-sumX(m1)*sumX(m1))*
                     (N*sumX2(m2)-sumX(m2)*sumX(m2)))
       end do
      end do
```

```
C Start evaluating higher order correlation coefficients
     do i = 2,7   !{ i = component subscript }
       call rank(i,mediator)
                               !{ order of correlation coefficient }
       do order = 2.6
         med = mediator(order) !{ subscript of the mediator }
         do n1 = 1, (med-1)
                              !{ subscript of the 1st variable }
           do n2 = (n1+1),7
                              !{ subscript of the 2nd variable }
             if ((n2 .1t. med) .or. (n2 .eq. i)) then
               a = corr(order-1,n1,n2)
               b = corr(order-1,n1,med)
               c = corr(order-1,n2,med)
               if (n2 .gt. med) c = corr(order-1, med, n2)
C -----
C To prevent negative sqrt argument and also divide by zero errors
C the correlation coefficient is set slightly less than 1 if
C either b or c >= 1.
               if (b.ge. 1.0) b = 0.999999
               if (c.ge. 1.0) c = 0.999999
               corr(order,n1,n2) = (a-b*c)/sqrt((1.0-b*b)*(1.0-c*c))
           end do
                     !{ end of n2 loop }
         end do
                     !{ end of n1 loop }
       end do
                     !{ end of order loop }
     end do
                     !{ end of i loop }
     do k = 1,6
        cpc(iR,k) = corr(6,1,k+1)
     end do
     return
     end
```

```
Subroutine Chart(scale, N)
C Construct and print the chart of the residuals of the difference
C between the estimated and measured values of R(i)
C
C
     Inputs: - scale = 10/resmax
С
                 N = no. of sets of data
C
    Outputs:- line printer plot of the chart of residuals written
С
              to logical unit number 4
С
С
    Associate routine :- Resetline(line)
     double precision res(6,200)
      character line(131), symbol(6)
      data symbol/'*','o','#','*','o','#'/
      common /residual/res,resmax
      write(4,5) resmax, (i, i=1,6)
  5 format(1H1,' Deviations of the measured R''s from'
     * 'the estimated R''s'/,' Maximum deviation = ',e12.5,//,
        6(10x, R', i1, 10x)
     do i = 1,N
       call resetline(line) !{ clear the line of characters }
       do j = 1,6
          index = res(j,i)*scale
          index = index + (j-1)*22 + 11
         line(index) = symbol(j)
        end do
       write(4,15) (line(k), k=1,131)
  15
       format(131a1)
      end do
     return
      end
```

```
Subroutine MatMult(m1,row1,col1,m2,row2,col2,product,error)
C Matrix multipication routine
C Input: m1, m2 (the first and second matrices)
       row1, col1, row2, col2 (dimensions of the two matrices)
C Output: product, error
С
C Error: 0 - successful operation
С
       1 - the column of the first matrix does not match that of !
C
          the second
integer row1, row2, col1, col2, error
    double precision m1(row1,col1),m2(row2,col2),product(row1,col2)
    if (col1 - row2) 10, 20, 10
 No. of columns in m1 does not match that in m2 - no operation
C-----
 10 error = 1
    return
C -----
 Perform the Matrix Multipications
C -----
 20 do i = 1, row1
      do j = 1, col2
       product(i,j) = 0.0
       do k = 1, col1
         product(i,j) = product(i,j) + m1(i,k)*m2(k,j)
       end do
      end do
    end do
    error = 0
    return
    end
```

```
Subroutine MatInv(dimen, matrix, inverse, error)
       Purpose: To calculate the inverse of a matrix
C
C
               with Gauss-Jordan elimination.
С
C
         Input : dimen, matrix
С
        Output : inverse, error
C
C
       Remarks: The original matrix will be changed to
C
                an identity matrix.
С
C
     Error code : 0 - no error
C
                1 - dimension of matrix < 1
С
                2 - no inverse exists
C
C
     Written by : S .Lam
С
          Date : 2 March 1988
C ------
     integer dimen, error, row, ref_row, term
     double precision matrix(dimen,dimen),inverse(dimen,dimen)
     double precision almostzero/1.0e-20/, multiplier
     common /zero/almostzero
     call Initial(dimen, matrix, inverse, error)
     if (dimen .le. 1) return
     ref_row = 0
     do while ((error .eq. 0) .and. (ref_row .lt. dimen))
      ref_row = ref_row + 1
C -----
   Check to see if the diagonal element is zero
C -----
       if (abs(matrix(ref_row,ref_row)) .lt. almostzero)
        call Pivot(dimen, ref_row, matrix, inverse, error)
       if (error) 30, 20, 30
C -----
  If error = 0 , carry out the inverse operation
C -----
        divisor = matrix(ref_row,ref_row)
  20
        do term = 1, dimen
          matrix(ref_row,term) = matrix(ref_row,term)/divisor
          inverse(ref_row,term) = inverse(ref_row,term)/divisor
        end do
```

```
do row = 1, dimen
C -----
  Make the ref_row element of this row zero
C -----
         if ((row .ne. ref_row) .and.
            (abs(matrix(row,ref_row)) .gt. a'mostzero)) then
           multiplier = -matrix(row,ref_row)/matrix(ref_row,ref_row)
           do term = 1, dimen
            matrix(row,term) = matrix(row,term) + multiplier*
                            matrix(ref_row,term)
             inverse(row,term) = inverse(row,term) + multiplier*
                            inverse(ref_row,term)
           end do
         end if
        end do
    If error <> 0, terminate operation
C -----
 30 end do !{ end of do while loop }
    return
    end
```

```
Subroutine Initial(dimen, matrix, inverse, error)
C -----!
  This subroutine tests for errors in the value of dimen !
С
  Input: dimen, matrix
  Output: inverse, error
C -----!
    integer dimen, error, row, col
    double precision matrix(dimen, dimen), inverse(dimen, dimen)
    double precision almostzero
    common /zero/almostzero
    error = 0
    if (dimen - 1) 10, 20, 30
C -----
   Case dimen < 1, error
C -----
 10 error = 1
    return
C ------
   Case dimen = 1, single element matrix
C -----
 20 continue
    if (abs(matrix(1,1)) .lt. almostzero) then
     error = 2 !{ singular matrix }
      inverse(1,1) = 1.0/matrix(1,1)
    end if
    return
C -----
    Case dimen > 1, make the inverse-to-be an identity matrix
 30 do col = 1, dimen
      do row = 1, dimen
       if (col .eq. row) then
         inverse(row, col) = 1.0
         inverse(row, col) = 0.0
       end if
      end do
    end do
    return
    end
```

```
Subroutine Pivot(dimen, ref_row, matrix, inverse, error)
C -----!
  This subroutine searches the ref_row column of the given matrix
  for the first non-zero element below the diagonal. If it finds
   one, then the subroutine switches rows so that the non-zero element !
  is on the diagonal. This same operation is applied to the inverse!
  matrix. If no non-zero element exists in a column, the matrix is
  singular and no inverse exists.
С
   Input: dimen, ref_row, matrix, inverse
  Output : matrix, inverse, error
C -----!
    integer dimen, error, new_row
    double precision matrix(dimen,dimen), inverse(dimen,dimen)
    double precision almostzero
    common /zero/almostzero
    error = 2 !{ No inverse exists }
    new_row = ref_row
C -----
  Try to find a row with a non-zero diagonal element
C -----
    do while ((error .gt. 0) .and. (new_row .lt. dimen))
      new_row = new_row + 1
      if (abs(matrix(new_row,ref_row)) .gt. almostzero) then
C -----
  Switch rows between the new_row and the ref_row
C -----
       call SwitchRow(matrix, dimen, new_row, ref_row)
       call SwitchRow(inverse, dimen, new_row, ref_row)
       error = 0
      end if
    end do !{ While }
    return
    end
```

```
Subroutine SwitchRow(matrix, dimen, row1, row2)
C ----!
С
   This subroutine switches the rows specified by row1 and row2 !
С
   of the given matrix.
С
C
   Input: matrix, dimen, row1, row2
С
   Output: matrix
C -----!
    integer dimen, row1, row2, term
    double precision matrix(dimen,dimen), dummy
    do term = 1, dimen
      dummy = matrix(row1,term)
      matrix(row1,term) = matrix(row2,term)
      matrix(row2,term) = dummy
    end do
    return
    end
```

6 Example

As a demonstration of the functions of the program SGBCalib, a "numerical" calibration was carried out using an "artificial" balance whose calibration coefficient matrix is given as follows:

	r 1	0	0	0	0	0	1
	0	1	0	0	0	0	
	0	0	1	0	0	0	
	0	0	0	1	0	0	
	0	0	0	0	1	0	
	0	0	0	0	0	1	
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	}
ĺ	0	0	0	0	0	0	
	0	0	0	0	0	0	
[C] =		0	0	0	0	0	
[C] =	0	0	0	0	0	0	. (18)
	0	0	0	0	0	0	, ,
	0 0 0 0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	1
	0	0	0	0	0	0	
	0	0	0	0	0	0	1
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	0	0	
	0	0	0	0	()	0	
	0	0	0	0	0	0	
	. 0	0	0	0	()	()	
	0	()	0	0	0	()	
	0	0	0	0	0	0	
ı	0	0	0	0	()	Ð	
	0	()	()	()	0	()]

This represents a perfectly linear balance with no interactions between the components. It is an ideal balance which, of course, is practically impossible to attain. Nevertheless, it serves the purpose of a simple illustration of the capabilities of the calibration program.

A number of sets of randomly generated component loads H_i , (ranged from 0 to 2 units) were used, in conjunction with the coefficient matrix [C] of equation (18), to produce a number of corresponding sets of strain gauge balance outputs R_i according to the relationship described by equation (2). The resulting R_i values were multiplied by a random factor in the range of ± 0.002 to simulate a random noise level of $\pm 0.2\%$ on the strain gauge balance outputs. The sets of H_i and noise inflicted R_i values were written to the input file SGBIN.DAT according to the format described in Section 4.2.1. The program was then executed to evaluate the elements of [C] from these sets of H_i and R_i values. The outputs of the program are shown in page 39 through to page 43.

Examination of the results shows that the maximum deviation of the estimated calibration coefficients is approximately 1% from the "true" values. This is hardly acceptable in the high precision work that a strain gauge balance is expected to operate. The inaccurate estimation of the coefficients in this particular case is because not all of the randomly applied component loads necessarily give useful independent information for the evaluation of the coefficients. The $\pm 0.2\%$ random noise level further distorts this information from accurately determining the coefficients. The accuracy, however, could be improved if a larger number of sets of loadings is used. Pages 44 and 45 are an abstract of the outputs from a similar test run except that 100 sets of random loadings were used for the "calibration". The results showed considerable improvement in the accuracies of estimation of the calibration coefficients.

On the other hand, the correlation coefficients show perfect multiple correlations between the strain gauge balance outputs and the six component loads. There is also perfect linear relationship between each of the R_i and its corresponding H_i . These are to be expected because the R_i values were derived from exactly the same relationship as the regression equation.

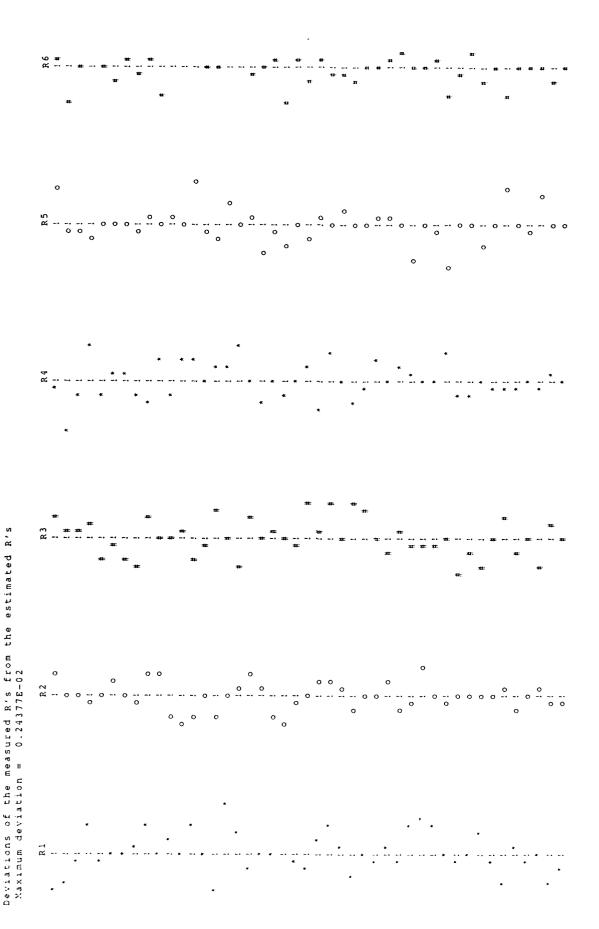
STRAIN GAUGE BALANCE CALIBRATION USING LEAST SQUARES METHOD

The Supplied Component Loads and Strain Gauge Bridge Readings:-

7 H			S.						1	o c
.4170	.1920	.6010	.6780	.5210	.9619	.4167	.1926	.6004	.6775	.5203
.2290	.6590	.8260	.8020	.5200	.0888	.2285	.6577	.8240	.8001	.5170
0.43700			0						0.02871	0.70307
.2620	.4790	.2310	.2650	.7840	.5971	.2604	.4797	.2326	.2625	.7855
.4330	.0060	.2440	.6040	.0790	.4935	.4327	.0065	.2451	.6049	.0781
.7330	.1800	.8590	.3920	.0110	.5259	.7313	.1801	.8582	.3912	.0100
.3050	.7100	.2250	.7860	.2400	.3427	.3051	.7105	. 2255	.7844	.2396
.7410	.9150	.2180	.4690	.9050	.2921	.7413	.9116	.2176	.4681	.9049
0565.	.7880	.4070	.6770	.0380	.9551	.4962	.7894	.4061	.6782	.0401
.6310	.2310	.3620	.1870	.6700	.9132	.6320	.2329	.3617	.1877	.6680
1.4810	.3600	.0680	.4350	.6550	.3884	.4787	.3615	.0676	.4352	.6542
.1800	.3030	.7410	.4090	.7620	.0013	.1800	.3028	.7408	.4102	.7622
.6710	.3750	.8210	.4780	.0430	.2015	6699.	.3747	.8209	.4804	.0443
1690	.0540	.4220	.8140	.0220	.5790	.1690	.0543	.4207	.8107	.0219
.6420	.6600	.0240	.5640	.7410	. 7342	.6412	.6610	.0241	.5638	.7402
0.3190	.1370	.6880	.4190	.1740	7945	.3196	1361	.6893	.4195	.1761
0.7760	.8490	.5660	.7540	.0680	.2757	.7762	.8469	.5683	.7546	0690.
1.6130	.8000	.0080	.9730	.7100	.4843	.6145	7967.	.0080	.9726	.7103
0.2660	.1770	.0230	.2390	.6870	.0297	.2659	.1772	.0211	.2385	.6852
.9820	.6260	.3890	.2240	.1380	.6569	.9792	6259	.3897	. 2234	.1380
0.4370	.1650	.1970	.0490	.5490	.0569	.4873	.1659	.1960	.0498	.5495
1.3760	.1450	.0950	.0670	.6160	.8014	.3759	.1447	.0953	.0666	.6176
0.7600	.1620	.0270	.6360	.2060	.0316	.7604	.1635	.0274	.6332	. 2043
1.2020	.3220	.7460	.5610	.9520	.1183	.2044	.3230	.7443	.5626	.9529
.4710	.7770	.8700	.3280	.5050	.8807	.4717	.7801	.8711	.3298	.5057
1420	.9220	.7890	.8250	.9860	.5771	.1426	.9225	.7877	.8235	.9826
.5830	.1770	.1650	.6640	.6550	.2929	.5833	.1789	.1640	.6650	.6545
.3460	.7450	.3490	.8060	.3430	.8859	.3463	.7471	.3482	.8066	.3435
.5100	.2630	.4640	.6000	.6660	.1426	.5123	.2644	.4651	.6008	.6684
.9930	.7430	.4500	.9270	.5600	.3249	.9952	.7396	.4502	.9282	.5594
.3720	.7200	.8480	.5890	.1650	.7331	.3727	.7200	.8471	.5885.	.1651
.2390	.0040	.6110	.6000	.5910	.4530	.2396	.0044	.6101	.5997	.5911
.6250	.9460	.7080	.5420	.5920	.9578	.6273	.9464	.7075	.5428	. 5924
.5300	.6220	.8210	.6290	.6890	.6842	.5312	.6202	.8215	.6290	.6881
.0670	8490	.9040	.6570	.7620	.2958	.0678	.8482	.9055	.6565	.7619
0.2930	.4280	.8860	.4720	.6540	.7373	. 2924	.4283	.8847	.4725	.6534
.2090	.5250	.3050	.1280	.8060	. 7449	.2099	. 5264	.3047	.1290	.8093
.0030	.6060	.1500	.3890	.4660	.0037	.0015	.6056	.1504	.3869	.4637
.0360	.8090	.6940	.7250	.4920	.5599	.0360	.8084	.6916	.7243	4919
.6250	.1840	.5970	.4630	.2900	.7005	.6253	.1844	. 5974	.4659	.2889
.1350	.6970	0590	.2520	.9120	.9033	.1359	.6938	.0586	.2504	.9151
.4310	.3940	.3160	.4910	.8510	.0086	.4300	.3944	.3160	.4907	.8504
.1980	.6520	.2450	.8800	.5510	.3677	.1994	.6517	. 2442	.8819	.5517
.1896	0606.	.3590	.7550	.1890	.9790	.1884	.9095	.3619	.7553	1887
0.8990	.7790	.9970	.3530	.7260	.3798	8995	.7786	.9956	.3529	.7245
,		,								

-0.0067772 0.0032534 0.0007855 0.0056942 0.0025058 0.00018483336 0.00084883336 0.00084883336 0.00084883336 0.00088483339 0.000884883336 0.000884883336 0.000884883336 0.000884883336 0.000884883336 0.000886893 0.000886893 0.00088884 .0041301 .0007379 .0019105 C20000133341
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000013341 .0031102 .0006310 .0037433 .0016760 coefficient calibration

Standard e	errors	of estimat	j					
R1 0.00.0	~# 9	R2 0.00107	R3 0.00135	3 135 0	R4 .00134	R5 0.00123	R6 0.0013	R6 0133
	υ (1	(
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000.1	00	1.00000	1.0000	0	1.00000	1.00000	1.0000	000
Coefficien	ក ១ ក	artial co	rrelation	<u>!</u>				
	1 2		R 2	83	ж 4	α	ır	9 8
1 :::	0000.	01		-0.32993	0.07504	-0.028	. ~	. ~
1	0.2	-	_	-0.20926	0.01755	0.0	8 10	0.13754
(۲)	0.3339	0	7363	1.00000	-0.00733	0.086	0	_
	0.1691	0		-0.07035	1.00000	. 13	4	
E S	0.1165	0	5475	0.01670	-0.06337	00	0	
up m	.1169	0		-0.10972	0.07715	.08	9 9	_
95% confide	ence le	vel of th	e linear	coeff.:-				
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001016	5.0	006635	0.0008364	0.0008300	00 0.0007651			
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.001391	. 0	007129	0 8 9	0.00089	_	O	18872	
. 30106.	m	006576	0.0008290	0.00082		0	08184	



An abstract of the outputs of a test case using 100 random sets of loadings

s,H	c1-	3	m	4	S	-95
0	.000171	.000933	.000429	.0000101	.000199	.000250
0	.000481	986666.	.000229	.000157	.000160	.000864
m	.000642	00195	999335	865000.	.000092	.0010
٠,	.000925	.000381	.000873	.999839	.000136	.000498
S	.000151	.000853	.000860	856000.	.999926	.000269
9	.000827	086000.	.000261	.001495	.000289	.000852
11	0.000697	.001510	.000275	.000669	.000372	.000
2.2	.000671	.000551	.000650	.000233	.000048	.000643
33	.000745	.000083	.001403	.001247	.00000.	.0000.
7.7	.001102	.000560	.000754	.000391	.000225	.000233
55	.00000	.001010	.000948	.000510	.000284	.0003
6 6	.000936	.000646	.000028	.001698	.000712	896000.
1.2	.000143	.000127	.000028	.000038	.000078	.0001
13	.000089	.000101	.000345	.000103	.000132	.000074
14	.000025	.000150	.000276	.00000.	.000296	.000174
15	.000116	.000201	.000092	160000.	.000109	.000129
16	.00003	.000132	.00001	.000031	.000124	.000175
23	.000052	860000.	.000143	.0000.	.000027	.00000.
2 4	.00000.	.000041	.000068	.000013	.000191	.000043
2.5	.000136	.000067	.000119	000000.	.000057	.000138
26	.000084	00182	.000050	.000082	.000038	.00026
ar M	.000068	.000045	.000225	.000174	690000.	.000037
3.5	.00015	00017	.000041	.000141	.000136	.00003
36	.0000.	.000013	.000031	.000064	.0000.	.000188
4.5	.000028	00020	050	.000025	600000.	00000.
9.5	.000046	.000084	.000103	.00000.	.000031	٠
9 9	.000030	.000012	.0000.	.000015	.000065	.0000.
	.00036	.000550	.000011	.000186	.000111	.000020
2	.000223	.000243	.000231	.000130	.000043	.0001
m	.000240	00022	.000483	.000435	.0000.	.000
-11 -21 -11	-0.0003904	-0.0001743	-0.0001214	-0.0000760	00	-0.0000446
ഗ	.00003	00322	0000	00000	721000	

	R6 0.00029		я6 1.00000		g Q	0.12863	-0.05072	-0.15832	0.03159	-0.04421	1.00000							
	ń		٥		u p	0.09371	-0.06523	-0.11368	0.00207	1.00000	-0.15263		- 9 -	0.0001040	2101000	6560000.0	0.3001033	7660006.0
	4 028 0.0002		R 0 1.00		Q	12							C5-	0.0000885		0.0000815	0.0000878	0.0000848
	84 9 0.000	ı	R4 0 1.0000		<u>~</u> ۵							coeff.:-	C4-	0.0001010	0.0000000	0.0000931	0.0001003	8960000.0
- : s e :	83 0.0002	correlation:	R3 1.00000	correlation:-	0			-0.03030 1				the linear co.	C3-	0.0001033	0.0001000	0.0000952	0.0001026	0660000.0
s of estimates:	R2 0.00031	f multiple	R2 1.00000	f partial	-							level of	C2-	0.0001107				
Standard errors	R1 0.00027	Coefficients o	R1 1.00000	Coefficients o	ρ	H1 1.00		H3 0.01407	•			95% confidence	C1-	0.0000983				

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16. ABSTRACT A computer program written in FORTRAN has been developed to calculate the calibration coefficients of a six-component strain gauge balance. The method of Least Squares is used in the calculation procedures which do not require the components of the balance to be loaded independently. The procedures are, in fact, independent of the way the balance is loaded as long as the supplied load/output data provide sufficient information for the determination of the coefficients. The program incorporates several statistical estimations for the evaluation of the performance of the

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16. ABSTRACT (CONT.) calibration process. It can th for the investigation of diffe to calibrate a strain gauge bal	rent loading	ed as an experimental tool schemes which may be used						
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